# $_1$  A kernel integral method to remove 2 biases in estimating trait turnover



- 22 Abstract
- 23

24 1. Trait diversity, including trait turnover, that differentiates the roles of species and 25 communities according to their functions, is a fundamental component of biodiversity. 26 Accurately capturing trait diversity is crucial to better understand and predict 27 community assembly, as well as the consequences of global change on community 28 resilience. Existing methods to compute trait turnover have limitations. Trait space 29 approaches based on minimum convex polygons only consider species with extreme 30 trait values. Tree-based approaches using dendrograms consider all species but distort 31 trait distance between species. More recent trait space methods using complex 32 polytopes try to harmonise the advantages of both methods, but their current 33 implementation have mathematical flaws.

34 2. We propose a new kernel integral method (KIM) to compute trait turnover, based on 35 the integration of kernel density estimators (KDEs) rather than using polytopes. We 36 explore how this difference and the computational aspects of the KDE computation can 37 influence the estimates of trait turnover. We compare our novel method to existing 38 ones using justified theoretical expectations for a large number of simulations in which 39 we control the number of species and the distribution of their traits. We illustrate the 40 practical application of KIM using plant species introduced to the Pacific Islands of 41 French Polynesia.

42 3. Analyses on simulated data show that KIM generates results better aligned with 43 theoretical expectations than other methods and is less sensitive to the total number of 44 species. Analyses for French Polynesia data also show that different methods can lead 45 to different conclusions about trait turnover, and that the choice of method should be 46 carefully considered based on the research question.

47 4. Mathematical aspects for computing trait turnover are crucial as they can have 48 important effects on the results and therefore lead to different conclusions. Our novel 49 kernel integral method generates values that better reflect the distribution of species in 50 the trait space than other existing methods. We therefore recommend using KIM in 51 future studies on trait turnover. In contrast, tree-based approaches should be kept for 52 phylogenetic diversity, as phylogenetic trees will then reflect the constrained speciation 53 process.

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- 56 Keywords: Beta diversity, Convex hull, French Polynesia, Trait turnover, Hypervolume,<br>57 Kernel, Traits
- Kernel, Traits

59 1. Introduction

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61 Biodiversity is a complex concept and can most easily be quantified by distinguishing three 62 complementary facets: taxonomic diversity based on a site-by-species matrix that captures 63 the compositional properties of a community; phylogenetic diversity that captures the 64 evolutionary relatedness among community members, using phylogenetic distance between 65 species alongside the site-by-species matrix; and trait diversity that describes a community 66 according to the traits of its residing species, using a species-by-trait matrix alongside the 67 site-by-species matrix (Devictor et al., 2010). The study of functional traits has been 68 advocated as fundamental to better understand and quantify community assembly (McGill 69 et al., 2006), as well as the impact of global change on community resilience and on the 70 ecosystem services that biodiversity provides (Gross et al., 2017). For example, through 71 comparison with null models and by relating traits to environmental gradients and to each 72 other, trait diversity can provide information about the assembly processes structuring an 73 ecological community (Ackerly & Cornwell, 2007), including biotic interactions between 74 species (Laureto et al., 2015). It also enables the estimation of components of ecosystem 75 function, such as nutrient use and storage, or ecosystem productivity (Cadotte et al., 2011; 76 Hillebrand & Matthiessen, 2009).

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78 In addition to the decomposition of biodiversity into taxonomic, trait and phylogenetic 79 components, unravelling how biodiversity is organised requires an understanding of how 80 assemblages of species are more or less similar to one another at different places and times, 81 i.e. turnover (Anderson et al., 2011). To do so, beta (β) diversity provides a direct link 82 between biodiversity at the regional (gamma  $-\gamma$  – diversity) and local (alpha –  $\alpha$  – diversity) 83 scales (Anderson et al., 2011; Chao et al., 2005, 2019). In particular, taxonomic β diversity 84 has been shown to be important for assessing the effects of conservation actions (Socolar et 85 al., 2016), for example for estimating the effect of the spatial distribution of protected areas 86 and their subdivision into multiple subareas on species diversity (Deane et al., 2022), or for 87 extrapolating regional species richness from limited data (Kunin et al., 2018). Although 88 having received less attention than taxonomic β diversity, trait turnover that describes 89 change in trait diversity across communities or regions has also been measured using β 90 diversity for similar applications (e.g. Carmona et al., 2012; Loiseau et al., 2017; Siefert et al.,

91 2013; Swenson et al., 2012; Villéger et al., 2013).

92

93 As a valuable and increasingly measured biodiversity facet, there are multiple important 94 steps to consider when estimating trait turnover over space or time. First, the precise choice 95 of traits can substantially influence the outcome (Petchey & Gaston, 2006). Second, despite 96 recent initiatives to collate large amounts of data for multiple traits across species (e.g. 97 Kattge et al., 2020; Middleton-Welling et al., 2020; Tobias et al., 2022), trait data are still 98 missing for many species and types of traits across taxonomic groups. Finally, and also the 99 focus of this work, different mathematical methods exist to compute trait diversity and 100 turnover that differ in outcome and therefore in the conclusions drawn in biodiversity 101 studies (Loiseau et al., 2017; Sobral et al., 2016; Villéger et al., 2017). A systematic 102 comparison can help identify an informative robust method and establish standards for 103 quantifying trait turnover.

104

105 There are two main categories of methods for calculating trait β diversity: (i) methods based 106 on the concept of trait space (referred to here as the 'trait space approach', and (ii) 107 methods that use dendrograms (referred to here as the 'tree-based approach'). The trait 108 space approach is based on a multi-dimensional space whose axes are determined by the 109 traits included in the analyses. Axes can correspond directly to the original traits or can be 110 derived from these traits through ordinations to reduce dimensionality. A particular species 111 typically represented as a single point in this trait space, and a polytope is computed as the 112 trait envelope of a set of points representing the species of a community or assemblage. The 113 minimum convex polytope (MCP), a convex hull, that encompasses all species of a 114 community in the trait space (Figure 1), has traditionally been used in these analyses 115 (Loiseau et al., 2017). As the MCP only captures information about the species with extreme 116 trait values in a community, it is sensitive to outliers and ignores how species are distributed 117 in the trait space, which can be crucial to delineate the functional roles of different species 118 within an ecosystem (Mouillot et al., 2021). Although other hull methods can be used to 119 compute the trait envelope (e.g. Irl et al., 2017), they are typically computationally intensive 120 and have been seldomly applied to  $\beta$  diversity analyses.

122 The tree-based approach consists of computing all pairwise distance between species based 123 on a set of traits, typically using the Gower distance to incorporate both continuous and 124 discrete traits. A clustering algorithm is then applied to these distances to generate a 125 dendrogram, from which measures of β diversity can be computed. Although the tree-based 126 approach considers all species in the computation of trait turnover, the dendrogram splits 127 into successive branches, and using the length of the branches connecting two species as a 128 measure of distance distorts the original trait distance between them compared to the 129 distance obtained through ordination in the trait space. In addition, the choice of the 130 clustering algorithm for generating the dendrogram will inevitably influence the outcome 131 (Loiseau et al., 2017).

132

133 The convex hull of trait space and the tree-based approach therefore make different 134 computational trade-offs, and the appropriateness of the two approaches for measuring 135 trait β diversity has been debated (Loiseau et al., 2017). In response to this debate and to 136 incorporate information from all species, Mammola & Cardoso (2020) proposed another 137 trait space approach where polytopes are defined by applying a threshold to the kernel 138 density estimation (KDE; Figure 1; see details in Methods below). The resulting polytope is 139 typically not convex, and its shape better reflects the distribution of species in the trait 140 space. Although it has the potential to provide a more accurate estimate of trait diversity 141 than the other two methods, this has not been formally assessed. The computational 142 aspects when computing kernel densities have largely been overlooked. These aspects, as 143 we plan to show here, are crucial so that all species contribute to  $\beta$  diversity in the 144 communities.

145

146 Here we propose a new trait space method, which we term the kernel integral method 147 (KIM), for computing trait  $\beta$  diversity based directly on the integration of the KDE rather 148 than on the polytope. We explore how the computational aspects of the KDE computation 149 can influence the estimates of trait  $\beta$  diversity with different methods. For comparison of 150 the existing and new methods, we use a set of theoretical examples for which we can justify 151 how trait β diversity metric should behave. We further apply the KIM method to compute 152 non-native plant trait turnover across islands and archipelagos of the Pacific Islands of 153 French Polynesia and compare results with the other methods.

154 155 2. Methods 156 157 2.1. Trait-space and tree-based approaches 158 159 2.1.1 Convex Hull 160

161 Computing trait turnover between two communities using the convex hull methods simply 162 consists in computing (i) the minimum convex polytopes (MCP) for each community, and (ii) 163 the hypervolumes of the intersection and the union of these two MCPs (Figure 1a,e). It is 164 then possible to compute a range of  $\beta$  diversity indices based on these four values. Here, 165 following Mammola & Cardoso (2020), we used the Jaccard dissimilarity index J (Jaccard, 166 1908) and the Williams replacement index W (Williams, 1996), defined as:

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J = 1 - \frac{A \cap B}{A \cup B} \tag{Eq.1}
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169 
$$
W = \frac{2 \times \min (A - A \cap B, B - A \cap B)}{A \cup B}
$$
 Eq.2

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171 where A and B are the hypervolumes of the MCPs for two communities. In our analyses, we 172 computed the MCPs and the indices using the hull.build() and hull.beta() 173 functions from the BAT R package V.2.8.1 (Cardoso et al., 2015, 2022). The Williams 174 replacement index evaluates the contribution of trait replacement to trait  $\beta$  diversity 175 (Carvalho et al., 2012, 2013), and the difference between Jaccard and Williams indices 176 quantifies how the trait richness difference between communities contributes to  $\beta$  diversity. 177 Although there are other approaches and indices that can decompose  $\beta$  diversity into 178 turnover replacement components, the relevance of these approaches is still debated 179 (Baselga, 2010; Baselga & Leprieur, 2015; Cardoso et al., 2014; Carvalho et al., 2012, 2013). 180 This debate is beyond the scope of this manuscript, and, to compare our methods, we only 181 followed the decomposition used by Mammola & Cardoso's (2020) (see sections on kernel 182 density hypervolumes below), readily available from the BAT R package (Cardoso et al., 183 2015, 2022).

185 The main issue with the convex hull methods is that it is insensitive to the addition or 186 removal of species within the MCP in the trait space (Figure 1). A corollary is that it is 187 sensitive to outliers, as they will define the MCP.

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189 2.1.2 Tree-based method

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191 The tree-based method consists in computing a dendrogram from the trait distance 192 between all species in the species pool (i.e. the entire list of species over all included sites, 193 not just those occurring in the pair of sites for each calculation of trait turnover; Figure 2a). 194 Multiple clustering algorithms can be used to generate the dendrogram, but here we 195 followed Loiseau et al. (2017) and used the unweighted pair group method with arithmetic 196 mean (UPGMA) algorithm, using the hclust() function from the stats R package (R Core 197 Team, 2022), as it has been shown to best conserve distances between species compared to 198 the original distances in the trait space.

199

200 For each site, a sub-tree including only the species present is generated by trimming the 201 overall tree (Figure 2b,c). It is then possible to compute the trees corresponding to the 202 union and the intersection of the two sub-trees (Figure 2d,e). We can then adapt Eqs 1 and 203 2 to compute the Jaccard and Williams indices, by using the total length of remaining 204 branches. Importantly, the sub-trees must be computed from the original tree generated 205 from the entire species pool, not those computed from only the residing species. This 206 conserves the internal branches in the union and the intersection of the two sub-trees, even 207 if these internal branches do not lead to any present species (see branch h in Figure 2e). 208 Therefore, for the example of Figure 2, Eqs 1 & 2 become:

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J = 1 - \frac{A \cap B}{A \cup B} = 1 - \frac{a+b+c+e+h}{a+b+c+e+f+g+h}
$$
 Eq.3

211 
$$
W = \frac{2 \times \min (A - A \cap B, B - A \cap B)}{A \cup B} = \frac{2 \times \min (f,g)}{a + b + c + e + f + g + h}
$$
 Eq.4

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213 The tree-based method therefore offers the advantage over the convex hull method that all 214 species will be accounted for when computing the  $\beta$  diversity indices. However, the 215 clustering algorithms often generate branch lengths between species in the dendrogram 216 that differ from the original distances in the trait space, which will necessarily influence the 217 value of any β diversity index (Loiseau et al., 2017).

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219 2.1.3. Kernel density hypervolumes (KDH)

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221 Mammola & Cardoso (2020) introduced the use of kernel density hypervolumes (KDH) for 222 computing indices of species turnover. This approach uses KDEs to generate polytopes that 223 are often non-convex (and can even be disjunct) and can be seen as a trait envelope around 224 the species points in the trait space. The recommended method is based on a Gaussian 225 estimator of the KDE (Mammola & Cardoso, 2020) and follows a series of four steps (see 226 Blonder et al., 2018 for further details): (i) points are drawn randomly within a hypersphere 227 around each species point in the trait space; (ii) these points are resampled to uniform 228 density; (iii) a KDE is computed from these points; (iv) a threshold (typically 95%) is applied 229 to truncate the KDE and define the polytope, from which hypervolumes can be computed. 230 The indices are then computed as per Eqs 1 & 2. In our analyses, we used the 231 kernel.beta() function from the BAT R package V.2.8.1 (Cardoso et al., 2015, 2022) to 232 apply the KDH method.

233

234 This method, although more computationally intensive than the convex hull method, allows 235 to account for the distribution of species points in the trait space to define the polytopes 236 and therefore the hypervolumes used in the computation of the turnover indices (Figure 1b-237 d, f-h). As a result, the KDH method is less sensitive to outliers.

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239 The KDH method nonetheless has some caveats. First, the choice of the threshold used to 240 construct the polytope will necessarily influence the components of the  $\beta$  diversity indices, 241 and therefore the final values. Second, by resampling random points to uniform density, 242 some information about the distribution of species points in the trait space is lost. Finally, in 243 the current implementation of the method in the BAT R package V.2.8.1 (Cardoso et al., 244 2015, 2022), the radius of the hyperspheres within which random points are drawn around 245 the species points and the bandwidth used during the computation of the KDE (the 246 bandwidth is a parameter that determines how smooth the KDE will be) are determined 247 based on the species point distribution of each community separately using the 248 estimate bandwidth() function from the hypervolume R package. As a result, the 249 more similar species are to each other, the closer random points will be to each other and 250 the KDE will show a steeper gradient (Figures A1-A24, see especially Figures A1, A9 and 251 A17). In other words, using different bandwidths and resampling random points to uniform 252 density gives different weights to a species depending on how different its traits are from 253 those of other species in the community. For a  $\beta$  diversity index to be unbiased we argue 254 that all species should have the same weight when relative abundance and intraspecific trait 255 variation are not concerned.

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- 257 2.2 A kernel integral method (KIM)
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259 Here we propose a novel computational method to compute trait turnover in the trait 260 space, to solve the issues associated with the KDH method. Our method computes  $β$ 261 diversity indices from the kernels themselves, therefore removing the influence of the 262 threshold used to generate the polytope, and uses different kernels than those used in the 263 KDH method. The KIM method consists of using only steps (i) and (iii) from the KDH method: 264 (i) points (typically 1000, but the number can be adjusted to account for species abundance, 265 for example) are drawn randomly within a hypersphere around each species point in the 266 trait space (the diameter of the hypersphere can be the same for all species, or reflect 267 intraspecific trait variability); (iii) a KDE is computed from these points, and rescaled 268 between [0,1]. From the KDE, we then propose the following equations to compute the 269 Jaccard dissimilarity index and the Williams replacement index:

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J = 1 - \frac{\int \min(\text{KDE}_A, \text{KDE}_B)}{\int \max(\text{KDE}_A, \text{KDE}_B)} \tag{Eq.5}
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$$
W = \frac{2 \times \min \left( \int KDE_A - \int \min \left( KDE_A, KDE_B \right) \int KDE_B - \int \min \left( KDE_A, KDE_B \right) \right)}{\int \max \left( KDE_A, KDE_B \right)}
$$
 Eq.6

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274 where KDEA and KDEB are the KDEs for communities A and B, and ∫KDEA is the integral of the 275 KDE for community A over all dimensions of the trait space. This is similar in essence to the 276 index of niche overlap proposed by Mouillot et al. (2005). In practice, since KDEs are 277 computed as multi-dimensional matrices, an integral is simply computed as the sum of all 278 elements of the matrix. The minimum and the maximum of two KDEs are analogue to the 279 intersection and the union of the polytope in the KDH method (Figure 3).

280

281 This kernel integral method enables us to overcome the limitations of the KDH method 282 mentioned above. First, there is no need to define a threshold: if the KDE is estimated over a 283 large enough area or volume, the local kernel density will approach zero and the integral 284 will therefore converge. Second, the radius within which the random points are drawn is the 285 same for all communities (but a suitable value must be chosen, which can be adjusted to 286 account for intraspecific trait variability). Finally, because there is no resampling to uniform 287 density, the distribution of species points in the trait space will be reflected more accurately 288 in the KDE.

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290 2.3. Test of the different methods on theoretical data and expectations

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292 We have described above the theoretical advantages and caveats of each of the different 293 methods. We implemented seven different methods in total to explore how the differences 294 between the characteristics of the methods can influence the results (Table 1):

- 295 A convex hull method (hereafter COVHULL).
- 296 A tree-based method (hereafter TREE).
- 297 The original kernel density hypervolume method with community-specific 298 bandwidths and uniform resampling (hereafter KDH V1).
- 299 A modified kernel density hypervolume method computed with the same bandwidth 300 for each pair of communities and uniform resampling (hereafter KDH V2), to explore 301 the influence of the bandwidth on the outcome.
- 302 A kernel integral method using kernel densities estimated with community-specific 303 bandwidths and uniform resampling (hereafter KIM V1), to explore the influence of 304 the kernel-based vs the polytope-based formulas (Figure 3).
- 305 A kernel integral method using kernel densities estimated with the same bandwidth 306 for both communities within a pair and uniform resampling (hereafter KIM V2), to 307 further explore the influence of the kernel-based vs the polytope-based formulas.

308 • A kernel integral method estimated with the same bandwidth for both communities 309 within a pair and without uniform resampling (hereafter KIM V3).

310

311 For each method, we computed Jaccard dissimilarity and Williams replacement, as defined 312 in Eqs 1-6. We then examined how these seven methods behaved in a set of theoretical 313 contexts for which we can make predictions of how an index of turnover should behave to 314 capture trait differences between communities.

315

316 In total, we simulated 72 different pairs of communities (Figures 4 and A1) and computed 317 our 14 indices (the Jaccard and Williams indices for each of the seven methods) for each 318 pair. For simplicity and computational efficiency, we used a trait space defined by two 319 theoretical traits. Each community was first delimited by a MCP represented by four species 320 arranged as a square. The MCPs were either of different sizes (square side of lengths 4 and 321 2; Figure 4) or of the same size (square side of length 4; Figure A1). They were either nested 322 within each other, partially overlapping, or disjunct. For each of these configurations, we 323 generated a community by randomly drawing species points within the MCPs according to 324 three patterns: (i) the species points were located in a small area (square sides of length 1) 325 in opposite corners of the MCPs (hereafter called the "different" point distribution); (ii) the 326 species points were randomly drawn within the MCPs (hereafter called the "random" point 327 distribution); (iii) the species points were located in a small area (squares of length 1) in the 328 closest corners of the MCPs (hereafter called the "similar" point distribution). We tested 329 these 18 configurations for 10, 40, 70 and 100 species points, and performed analyses 50 330 times for each of the resulting 72 configurations (2 MCP size setups x 3 relative positions x 3 331 random point distributions x 4 sets of point numbers x 50 repeats = 72 configurations x 50 332 repeats).

333

334 There is one obvious difference between these theoretical communities and communities 335 that would be analysed for real-world applications: real-world communities belonging to the 336 same ecological system such as those described in the next section will usually share 337 species, resulting in many species points overlapping in the trait space. Here we used 338 independent random species distributions in the trait space for the two communities in 339 order to have more flexibility in these species distributions, to explore in detail how each of 340 the seven methods would behave across a wide variety of extreme configurations, and 341 better test disentangle the implications of their computational specificities.

342

343 This flexibility allows us to describe how a  $\beta$  diversity index should behave based on what it 344 is supposed to capture from these theoretical configurations. These expectations are 345 depicted in Figures 5, A2, A3, and their justification provided in Table A1. In summary, 346 Jaccard dissimilarity should increase as most species points in the two communities move 347 away from each other. The replacement component should decrease if the difference in 348 area covered by the two sets of species points increases. These patterns should be 349 especially clear for large numbers of species, i.e. for high densities of species points. For few 350 species and low species point density (e.g. when only 10 species points were randomly 351 drawn in the trait space), we expect these patterns to be weak, as the stochastic element of 352 the species point distributions may obscure the results.

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354 2.4. Established non-native plants in French Polynesia

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356 To examine how the different methods may lead to different conclusions when analysing 357 real data, we examined the trait diversity of plant species introduced to the Pacific islands of 358 French Polynesia, comparing trait turnover across islands and archipelagos using each 359 method. We extracted data from PacIFlora (Wohlwend et al., 2021). For French Polynesia, 360 PacIFlora contains data on naturalised non-native plant species across the 80 Pacific islands 361 over five archipelagos: The Society Islands, the Gambier Islands, the Tuamotu Islands, the 362 Tubuai Islands, and the Marquesas. However, careful examination of the database revealed 363 that some species were not naturalised but cultivated or endemic. We therefore only used 364 the 417 naturalised species in PaciFlora appearing in the Appendix of Fourdrigniez & Meyer 365 (2008).

366

367 For these 417 species, data on species woodiness (woody vs. herbaceous species), seed 368 mass, plant height and specific leaf area (SLA) were extracted from multiple trait databases, 369 including TRY (Kattge et al., 2011, 2020), LEDA (Kleyer et al., 2008), PLANTATT (Hill et al., 370 2004), Austraits (Falster et al., 2021), BIEN (Maitner, 2022), EcoFlora (Fitter & Peat, 1994), 371 and BROT (Tavşanoğlu & Pausas, 2018). Seed mass, plant height and SLA have been used to 372 characterise different plant life strategies (Díaz et al., 2016; Westoby, 1998). When different 373 databases contained different values, we used the mean for seed mass, plant height and 374 SLA, and the most frequent category for woodiness. Data on plant woodiness was available 375 for all 417 species. Trait data for seed mass and plant height were only available for 250 out 376 of the 417 species. Data for seed mass, plant height and SLA were only available for 124 out 377 of 417 species. We therefore performed three sets of analyses: (i) a set for the 250 species 378 with data on seed mass and plant height, (ii) a set for the 124 species with data on the three 379 traits, and (iii) a set for the same 124 species, using data on seed mass and plant height only, 380 to assess the robustness of the results to data availability and trait selection. In the 381 following we present and discuss mainly results for seed mass and plant height for the 250 382 species (see Figure D1 for the distribution of plant species in this two-dimensional trait 383 space), as it represents more than half of the species and should be less biased despite using 384 only two traits.

385

386 Prior to analysis, seed mass, plant height and SLA were log-transformed and rescaled 387 between [0,1], so that the traits would be more uniformly distributed in the trait space. We 388 then computed the Jaccard dissimilarity and the Williams replacement indices for all species 389 together, and for woody and herbaceous species separately, to have a more comprehensive 390 assessment of potential differences between methods, as woody and herbaceous plants 391 tend to characterise different parts of the global spectrum of plant forms and functions 392 (Díaz et al., 2016). We also computed these indices for all French Polynesian islands, and for 393 each archipelago separately.

394

395 Finally, the behaviours of the different indices were analysed using randomisation tests. We 396 randomised the presence-absence matrices for all islands and for each archipelago by 397 keeping species occupancy and island richness constant (i.e. the sim9 algorithm from 398 (Gotelli, 2000)), and compared the Jaccard dissimilarity and Williams replacement indices 399 generated by the 7 methods for the original matrices to the indices computed over 10 400 randomised matrices for each original matrix (the number of randomisations was 401 constrained by computation time).

402

403 The purpose of these analyses on real data was only to examine how results may differ 404 between methods for more complex data than used in the theoretical analyses, to assess 405 each method's range of sensitivity. As each archipelago contains multiple different islands 406 whose combinations will fall across a large spectrum of trait profile configurations, it was 407 not possible to define a priori expectations and the purpose is therefore not to determine 408 which methods are in line or not with a priori expectations, contrary to the theoretical 409 analyses.

- 410
- 411 3. Results
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413 3.1. Theoretical data

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415 Overall, KDH and KIM tended to converge towards similar values and behaviours as the 416 number of species points increased (Figures 6-8, C1-C12), corresponding to theoretical 417 expectations (Figures 5, A2, A3). In contrast, the convex hull and the tree-based methods 418 generated indices of turnover different from the other methods and from the theoretical 419 expectations. The main differences between observed and expected values for all methods, 420 except the tree-based method, were for the contribution of replacement to overall turnover 421 (computed as the ratio of the Williams replacement index to the Jaccard dissimilarity index) 422 for MCPs of the same size in the nested / random and the nested / similar configurations, 423 which was lower than the expected value of 1 (Figures A3, C6). This is likely due to the fact 424 these are the configurations for which the values of the Jaccard index are small, and small 425 changes in Williams replacement index due to stochasticity in the distribution of species 426 points in the trait space will be disproportionally large.

427

428 For all indices of turnover, the three KIM methods generated values above 0.5 and above 429 other methods when the point distributions were different from each other (i.e. the 430 "Different" point distributions under all MCP configurations, and for all three point 431 distributions under the "Disjunct" MCP configuration). KIM V3 generated values below 0.5 432 and below other methods when the point distributions were similar from each other (i.e. 433 the "Similar" point distributions under all MCP configurations), and intermediate values

434 otherwise, in-between the values generated by the other methods. These results suggest 435 KIM V3 can better distinguish between different species point distributions in the trait space 436 (Figures 6-8, C1-C12). The KIM V3 method also tended to be less sensitive to the number of 437 species points than the other KDH and KIM methods, with values and behaviours being 438 similar from 10 to 100 species points.

439

440 When communities had MCPs of the same size, the KIM V1 and V2 methods generated 441 similar results to the KDH V1 and V2 methods, respectively, for all β diversity indices 442 (Figures C4-C6, C10-C12). However, when the MCPs had different sizes, the KIM methods 443 tended to generate values more similar to each other than to the KDH methods (Figures 6-8, 444 C1-C3, C7-C8).

445

446 Adjusting the bandwidth to be common between communities in each pair in the 447 computation of the kernels for the KDH and KIM methods (i.e. switching from V1 to V2) 448 resulted in lower dissimilarity values, both for the Jaccard dissimilarity index and the 449 Williams replacement index, for all configurations. This is because the radius of the 450 hyperspheres and therefore the steepness of the kernels were the same for both 451 communities in the V2 methods, increasing similarity. The effect of removing the resampling 452 of random points to uniform density (i.e. from KIM V2 to KIM V3) had an often larger and 453 more variable effect than adjusting the bandwidth, as the values generated by KIM V3 could 454 be either larger, smaller or in between those of the KIM V1 and V2 methods.

455

456 3.2. Established non-native plants in French Polynesia

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458 Raw values of Jaccard dissimilarity, of Williams replacement and of the contribution of 459 replacement to turnover differed greatly between the different methods. Maximum 460 differences in values between methods were around 0.6 for Jaccard dissimilarity, 0.2 for 461 Williams replacement, and 0.8 for the contribution of replacement to turnover (Figures 9, 462 C2, C3). The KIM V3 and the KDH V2 methods generated the lowest Jaccard dissimilarity, 463 and KIM V1 and TREE the highest. In contrast, KIM V3 consistently generated much higher 464 values for the contribution of replacement to turnover than other methods, as expected 465 from the fact that it better accounts for differences in species point distributions in the trait

466 space. Results showed similar trends for all the combinations of traits and number of 467 species used in the analyses (Figures D2, D3).

468

469 Importantly, compared to the other methods, the KIM methods sometimes generated a 470 different ranking between archipelagos for Jaccard dissimilarity. This is especially true for 471 woody species, for which KIM V3 suggests that trait turnover was higher for the Gambier 472 than for any other archipelagos, for all combinations of traits, and for both Jaccard and 473 Williams replacement indices (Figures 9, D2, D3). By contrast, the other methods generated 474 results that were more variable depending on the combination of traits and species used.

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476 Randomisation of presence-absence matrices show that the KDH V2, KIM V2, KIM V3 and 477 Tree methods tended to generate more consistent values for Jaccard dissimilarity compared 478 to the convex hull, KDH V1 and KIM V1 methods (Figure D4). For Williams replacement, 479 values were also more consistent across randomised matrices for KIM V3 than for the other 480 methods, especially for herbaceous species.

481

482 4. Discussion

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484 Here we compared existing and novel methods to compute trait turnover for simulated and 485 empirical data, to illustrate how differences in the computational aspects of these methods 486 reflect different aspects of trait diversity and can affect inferences made from trait diversity 487 comparisons.

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489 4.1. Theoretical aspects of trait turnover computation

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491 Comparing the seven methods using simulated data, for which we had complete control of 492 the community trait profiles in the trait space, revealed the important effects of the 493 computational specifics of each method on the value of trait  $\beta$  diversity. In particular, we 494 assessed the effect of conserving trait distance between species by comparing the tree-495 based method, which distorts the trait distance between species in the dendrogram (Maire 496 et al., 2015), to the trait space-based methods. The tree-based method consistently 497 generated results most different from theoretical expectations (Figures 5-8). It tended to 498 either underestimate or overestimate dissimilarity in the trait profiles of communities in the 499 MCPs, and tended to generate high values for the Williams replacement index.

500

501 Interestingly, the convex-hull methods produced results similar to the tree-based method 502 for Jaccard dissimilarity, also generating results that differed from theoretical expectations. 503 This is because CONVHULL only uses a subset of the species in the trait space. Although 504 using a trait space approach better conserves trait distance between species than a tree-505 based approach when all species are included, this property is broken when species are 506 ignored in the computation of trait turnover. Therefore, although the CONVHULL and tree-507 based methods have been contrasted in the literature and have been shown to generate 508 different results (e.g. Loiseau et al., 2017), we show that neither of these two methods 509 accurately reflects the trait profile of a community.

510

511 In contrast, the other five trait space methods compared in this article (KDH V1-V2 and KIM 512 V1-V3) generated results more in line with theoretical expectations. These methods 513 therefore offer a more consistent representation of the community trait profile in the trait 514 space, i.e. they better capture the contribution of all species to the assessment of turnover. 515 The computational aspects of these approaches to estimating trait turnover have 516 nonetheless important effects on the generated β values, with potential consequences for 517 inferences about made trait turnover in an assemblage or community.

518

519 Specifically, we explored three computational aspects of these methods: (i) the use of 520 polytopes vs kernel integrals (KDH V1 vs KIM V1; Eqs 1 & 2 vs Eqs 5 & 6); (ii) the use of the 521 same or different bandwidths for each community in a pair when computing the KDE (KDH 522 V1 vs V2 and KIM V1 vs V2); and (iii) the use of point resampling when computing the KDE 523 (KIM V2 vs V3). All three aspects proved to have important effects on the β diversity values 524 calculated. Using kernel integrals, the same bandwidth and not resampling (i.e. using KIM 525 V3) generated results most in line with theoretical expectations.

526

527 The respective effects of these three computational aspects on trait turnover (β values) 528 depend on the index used (Jaccard dissimilarity or Williams replacement) and on the 529 configuration of the community trait profiles. For example, Jaccard dissimilarity is sensitive 530 to the difference in bandwidth between communities (Figure 6). This is because using 531 different bandwidths changes the shape of the KDEs (akin to making the distributions larger 532 or narrower in Figure 3) and generates polytopes with different areas (akin to changing the 533 lengths of A and B in Figure 3). Consequently, Jaccard dissimilarity values reflect this 534 artificial difference in trait richness, but not by Williams replacement. In contrast, for 535 Williams replacement, the use of polytopes or kernel integrals proved to be the most 536 important factor (Figure 7). This is because kernel integrals better reflect small variations in 537 the shape of the KDE (akin to changing the shape of the distributions and the overlapping 538 area in Figure 3) and thus better capture trait replacement. Similarly, resampling also 539 affected Williams replacement for communities with an "overlapping" configuration (Figure 540 7), especially for species-poor communities. This is because, when compared to the more 541 uniformly distributed trait profile of species-rich communities, each species has a greater 542 weight on the shape of the trait profile in species-poor communities, and the idiosyncrasy in 543 the position of different species can drastically change trait profiles if resampling is not 544 applied.

545

546 Both trait envelope, as captured by a convex hull, and kernel-based community trait profiles 547 have complementary uses, and the choice of an analytical approach will depend on the 548 research or management question. On the one hand, species with extreme trait values 549 defining a trait envelope for a given community can help capture the whole range of trait 550 values of species that may potentially join the community. The trait envelope may therefore 551 be an important piece of information to assess the risk of potential invaders to a region ( 552 e.g. the join the locals vs. the try harder hypothesis; Tecco et al., 2010), or an indicator of 553 the loss of trait extremes. The CONVHULL method is appropriate for such applications. On 554 the other hand, capturing the trait distribution of all species in a community in the trait 555 space provides a more comprehensive description of trait diversity and is necessary for 556 identifying community assembly processes (Falster et al., 2017). The distributional profile of 557 species in the trait space can also highlight gaps within the trait envelope, where introduced 558 species with corresponding traits could have a higher chance to establish (i.e. the "empty 559 niche hypothesis"; MacArthur, 1970; Molofsky et al., 2022). This community trait profile 560 thus reflects ecosystem resilience and trait redundancy (Hui et al., 2021; Mouillot et al.,

561 2021). Our results suggest that the KIM V3 method is most informative and least biased for 562 addressing trait diversity questions.

563

564 4.2. Empirical test of methods using plant data from French Polynesia

565

566 The empirical testing of these methods provided further insight on the behaviour of the 567 different methods for a mixture of configurations of species points in the trait space (Figure 568 D1), and on how using different indices can lead to different conclusions. The KIM V3 569 method consistently generated much higher values for the contribution of replacement to 570 turnover than other methods, even when randomising site-by-species matrices (Figures 9, 571 D2-D4), suggesting that the higher Jaccard dissimilarity values generated by the other 572 methods may reflect an overestimation of the contribution of trait richness difference (the 573 complement of replacement) to turnover. Importantly, depending on the method used, one 574 could either conclude that most islands of an archipelago are very different from each other 575 in terms of trait diversity (e.g. Jaccard dissimilarity values > 0.5 for KIM V1 for the Gambier, 576 Tuamotu and Society archipelagos), or very similar (Jaccard dissimilarity values mostly < 0.2 577 for KIM V3). These different conclusions, in addition to the different rankings generated by 578 the different methods could be crucial for conservation decisions. For example, assuming 579 management actions are influenced by species traits, low trait dissimilarity between islands, 580 as indicated by KIM V3, would suggest that a similar management approach is appropriate 581 across most islands, simplifying management and potentially improving management 582 efficiency. In addition, the high contribution of replacement to turnover suggests that 583 existing differences in community trait profiles are unlikely to be the result of differences in 584 colonisation pressure, and may point towards either idiosyncratic or niche-driven factors.

585

## 586 5. Final recommendations / Conclusion

587

588 The kernel integral method presented here computes trait β diversity by directly integrating 589 KDEs. Out of the three different indices this method can generate, KIM V3 implements the 590 same bandwidth for the paired communities without resampling random points to uniform 591 distribution. KIM V3 generates values that better reflect the distribution of species in the 592 trait space (i.e. the community trait profiles) than methods based on convex hulls or 593 dendrograms, and also better than other methods based on KDEs. The approach is also 594 flexible, information rich and readily adapted to account for relative abundance between 595 species and intraspecific trait variation, by using different numbers of random points and 596 radii to generate the KDEs. Together with the convex hull method to inform on the trait 597 envelope, and tree-based approaches for quantifying phylogenetic diversity, and the kernel 598 integral method using the same bandwidth and non-uniform point distribution provide a 599 complementary set of metrics for understanding patterns of trait diversity and turnover. 600

601

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# 817 Tables

818



# 819 Table 1. Summary of the methods used to compute trait turnover.





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826 827 Figure 1. Summary of the trait space approaches for two communities with different species<br>827 Figure 1. Summary of the trait space approaches for two communities with different species<br>828 Fin a two-dimensional tr 829 additional species present in community 2, resulting in the same outcome when computing 830  $\beta$  diversity metrics. b, f) The KDH (kernel density hypervolume) method generates a polytope 831 for each community, whose shape will vary with the absence or presence of the additional 832 species in community 2 and is often non-convex. As a result, the outcome of the Jaccard 832 species in community 2 and is often non-convex. As a result, the outcome of the Jaccard 833 dissimilarity or the Williams replacement formulas will differ. c,d) KDEs corresponding to the 834 polytopes in (b). g, h) KDEs corresponding to the polytopes in (f). 835



836<br>837 Figure 2. Components of the tree-based approach for the computation of trait turnover<br>838 between two sites whose species are part of a larger species pool. a) Dendrogram for all 839 species in the species pool. b) Dendrogram for species present in site A, after trimming the 840 global tree from a). c) Dendrogram for species present in site B. d) Dendrogram for species 841 gresent in either site A or site B. Since species 3 does not occur at any of the two sites, 841 present in either site A or site B. Since species 3 does not occur at any of the two sites, 842 branch d is not included. e) Dendrogram for species present in both sites A and B. Although 843 species 4 and 5 are present in only one of the two sites, branch h appears in both<br>844 dendrograms, and is therefore conserved. dendrograms, and is therefore conserved. 845



846<br>847<br>848 848 Figure 3. Details of the computation of trait turnover for two pairs of communities (one pair<br>849 in (a) and one in (b)) following different approaches. In each graph (a,b), the curves fictional 850 KDEs (Kernel density estimators) for the two communities, in one dimension (for 851 simplification, we assume their densities are 0 beyond intersecting the horizontal axis). 852 Using the KIM (Kernel integration method) formula, The Jaccard index is computed as one 853 minus the area in grey divided by the striped area (Eq. 5), and the value is different for the

854 two pairs of communities in (a) and (b) (as is the value of Willams replacement index, Eq. 6).

855 (c) The horizontal segments represent one-dimensional polytopes (defined using the values

856 where the KDEs intersect the horizontal axis for simplification), used to compute the Jaccard

857 or Williams indices using the KDH (kernel density hypervolume) method (Eqs 1, 2). Contrary 857 or Williams indices using the KDH (kernel density hypervolume) method (Eqs 1, 2). Contrary

858 to KIM, the KDH method only generates a single value for each index for the two pairs of 859 communities.





861<br>862<br>863

863 Figure 4. One of the 50 instances of the nine theoretical configurations of pairs of<br>864 communities for different sizes of the MCPs (see Figure A1 for same size MCPs), using only

865 10 randomly drawn species points for clarity (point distributions for 40, 70 and 100 species

866 were also generated). We varied how the MCPs overlapped ("nested", "overlapping" or

867 "disjunct"), and how the species points are distributed within the MCPs ("different" -

868 distributed in opposite corners of the MCPs -, "random" – randomly distributed within the 868 distributed in opposite corners of the MCPs –, "random" – randomly distributed within the

869 – MCPs – or "similar" – distributed either within the same small area, or in the closest corners.<br>870 – Julie MCP-1 870 of the MCPs).





 $\frac{8}{3}$ 

875 Figure 5. Qualitative differences in β values for trait turnover predicted under different<br>876 Simulated configurations of species points in the trait space when the MCPs of the two 877 communities have different sizes. Jaccard dissimilarity (squares, solid lines), Williams 878 replacement (triangles, dotted lines), and contribution of replacement to overall turnover, 879 computed as Williams replacement divided by Jaccard dissimilarity (circles, dashed lines). 880 Cverall,  $\beta$  values, are expected to decrease as the overlap between the trait profiles of the 881 two simulated communities increases. Jaccard dissimilarity accounts for differences in the 882 spread of the trait profiles in the trait space (i.e. trait richness), whereas Williams 882 spread of the trait profiles in the trait space (i.e. trait richness), whereas Williams 883 replacement is independent from differences in trait richness. For detailed explanations<br>884 replacement is above as in 0 values, see Table 44 in Cl 884 about the changes in  $\beta$  values, see Table A1 in SI.<br>885 885<br>886



889 Table 1, for MCPs of different sizes, for 10 and 100 species points (see Appendix C for the 890 full set of results). 890 full set of results).



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895 full set of results).





899 Williams replacement divided by Jaccard dissimilarity, for the seven methods summarised in the seven methods summarised in

900 Table 1, for MCPs of different sizes (see Appendix C for the full set of results).



902<br>903<br>904 904 Figure 9. Application of seven trait turnover methods to empirical data on French<br>905 Polynesian plant species. Jaccard dissimilarity index, Williams replacement index, and 906 contribution of replacement to overall turnover, computed as Williams replacement divided 907 by Jaccard dissimilarity, for the seven methods summarised in Table 1, for French Polynesia 908 (FP) and its archipelagos, for all species, woody species and herbaceous species, using seed 908 (FP) and its archipelagos, for all species, woody species and herbaceous species, using seed 909 mass and plant height, for 250 out of 417 species (see Figures C2 and C3 for using seed 910 mass, plant height and SLA on 124 species, and seed mass and plant height on 124 species).<br>911 Note that the order of the islands is arbitrary, and lines between symbols are used as a 911 Note that the order of the islands is arbitrary, and lines between symbols are used as a<br>912 visual aid and not to depict continuous change. visual aid and not to depict continuous change.